

## Self-similarity of Brownian motion and a large deviation principle for random fields on a binary tree

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Received September 17, 1991; in revised form June 30, 1993

**Summary.** Using self-similarity of Brownian motion and its representation as a product measure on a binary tree, we construct a random sequence of probability measures which converges to the distribution of the Brownian bridge. We establish a large deviation principle for random fields on a binary tree. This leads to a class of probability measures with a certain self-similarity property. The same construction can be carried out for  $C[0, 1]$ -valued processes and we can describe, for instance, a  $C[0, 1]$ -valued Ornstein–Uhlenbeck process as a large deviation of Brownian sheet.

*Mathematics Subject Classifications (1991):* 60F10, 60J65

### Introduction

Self-similarity of Brownian motion induces a certain ergodic behaviour of the Brownian bridge. We investigate large deviations of this ergodic behaviour. Let  $C[0, 1]_{0,0}$  be the space of all functions  $X$  in  $C[0, 1]$  with  $X(0) = X(1) = 0$  and  $P$  the distribution of the Brownian bridge. We define mappings  $T_0, T_1$  of  $C[0, 1]_{0,0}$  on itself which describe rescalings of the left and the right half of the function  $X: (T_0 X)_t = \sqrt{2}(X_{t/2} - tX_{1/2}), (T_1 X)_t = \sqrt{2}(X_{(t+1)/2} - (1-t)X_{1/2})$ , respectively. Due to the self-similarity of Brownian motion,  $P$  is invariant and, in fact, even ergodic under  $T_0$  and  $T_1$ .

For each function  $\omega \in C[0, 1]_{0,0}$  and each  $\theta \in \{0, 1\}^{\mathbb{N}}$ , we now construct a sequence of probability distributions  $R_{n,\theta}(\omega)$  on  $C[0, 1]_{0,0}: R_{n,\theta}(\omega) = (1/n) \sum_{k=0}^{n-1} \delta_{T_{\theta_k} \dots T_{\theta_0} \omega}$ , where  $\delta_\omega$  denotes Dirac measure on  $\omega$ . This means that in each step we choose, according to  $\theta$ , the left or the right half of the function  $\omega$  and rescale it.  $R_{n,\theta}(\omega)$  is the empirical distribution corresponding to this sequence of functions. Let  $\theta_1, \theta_2, \dots$  be independent coin tossings under  $\lambda$ . We can show that  $R_{n,\theta}(\omega)$  converges to  $P$  for  $P$ -a.e.  $\omega$  and  $\lambda$ -a.e. all  $\theta$ . This ergodic behaviour of the Brownian bridge  $P$  says that we can reconstruct  $P$  with an “infinitesimal” piece of a single “typical” trajectory around a “typical” point of the unit interval  $[0, 1]$ , if we identify  $\lambda$  with Lebesgue measure.

We get another description of  $T_0$  and  $T_1$  using the Lévy–Ciesielski construction of the Brownian bridge: each function in  $C[0, 1]_{0,0}$  can be written as a superposition of the Schauder functions  $e_{n,k}$ ,  $k = 1, 2, \dots, 2^{n-1}$ ,  $n = 1, 2, \dots$ :

$X(t) = \sum_{n=1}^{\infty} \sum_{k=1}^{2^{n-1}} Y_{n,k}(X) e_{n,k}(t)$ ,  $0 \leq t \leq 1$ . This defines a mapping of  $C[0, 1]_{0,0}$  into  $\mathbb{R}^I$ , where  $I := \{(n, k) | k = 1, 2, \dots, 2^{n-1}, n = 1, 2, \dots\}$  has the structure of a binary tree.  $T_0$  and  $T_1$ , interpreted as mappings of  $\mathbb{R}^I$  into  $\mathbb{R}^I$ , correspond to shifts of the tree to the left and to the right, respectively.  $P$  corresponds to the product measure on  $\mathbb{R}^I$  with marginal distribution  $N(0, 1)$ , i.e. the random variables  $Y_{n,k}$ ,  $k = 1, 2, \dots, 2^{n-1}$ ,  $n = 1, 2, \dots$  are independent with distribution  $N(0, 1)$  under  $P$ . Similar representations of stochastic processes as tree-indexed random variables have been investigated recently in the context of wavelet transforms, see [2].

We now look at large deviations of the convergence of  $R_{n,\theta}$  to  $P$ . Note that  $R_{n,\theta}$  would not correspond, on the lattice  $\mathbb{Z}^d$ , to the usual empirical field, but rather to the empirical distribution of a sequence of configurations we get shifting along the path of a (transient) random walk on the lattice. The same model has been considered independently by Ben Arous and Tamura. They get, for each fixed  $\theta$ , a large deviation principle for the distributions of  $R_{n,\theta}$ , where the rate function depends on  $\theta$ . We are interested in “uniform” bounds; in particular, we want to look at probability measures on the tree which are invariant under all shifts, not only for a fixed  $\theta$ . We prove that the finite-dimensional marginals of  $R_{n,\theta}$  satisfy a large deviation principle and characterize the rate function as a “mean entropy” (Theorem 4.1). Minimizing this rate function leads to the class of *self-similar* probability measures, defined by invariance under  $T_0$  and  $T_1$ . Such a self-similar probability distribution on  $\mathbb{R}^I$  can be identified, under certain conditions, with a probability distribution on  $C[0, 1]_{0,0}$ . We investigate some properties of the corresponding “self-similar” stochastic processes.

More generally, we may replace  $\mathbb{R}$  with a Polish space  $S$  and  $P$  with a product measure on  $S^I$ . If we set  $S = C[0, 1]$  (see Sect. 5), the product measure on  $S^I$  with Wiener measure as one-dimensional marginal can be identified with the distribution of Brownian sheet. We can describe then, for instance, a  $C[0, 1]$ -valued Ornstein–Uhlenbeck process as a large deviation of Brownian sheet.

## 1 Lévy representation of functions in $C[0, 1]_0$ as elements of $\mathbb{R}^I_0$

We consider the following representation of functions in  $C[0, 1]_0$  which is the space of all functions  $X$  in  $C[0, 1]$  with  $X(0) = 0$ . Let the *Haar functions*  $\varphi_0$ ,  $\varphi_{n,k}$ ,  $k = 1, 2, \dots, 2^{n-1}$ ,  $n = 1, 2, \dots$  be defined as

$$\varphi_0(t) = 1 \quad 0 \leq t \leq 1$$

and

$$\varphi_{n,k}(t) = \begin{cases} 2^{(n-1)/2} & (k-1)/2^{n-1} \leq t < (k-1/2)/2^{n-1} \\ -2^{(n-1)/2} & (k-1/2)/2^{n-1} \leq t < k/2^{n-1} \\ 0 & \text{else.} \end{cases}$$

Then  $\{\varphi_0, \varphi_{n,k}, k = 1, 2, \dots, 2^{n-1}, n = 1, 2, \dots\}$  forms a complete, orthonormal system in  $L^2[0, 1]$ . It is the oldest example of an orthonormal wavelet basis with

“mother wavelet”  $\varphi_{1,1}$  (see, for instance, [4]). We won't represent  $X$  as a superposition of  $\varphi_0, \varphi_{n,k}, k = 1, 2, \dots, 2^{n-1}, n = 1, 2, \dots$  (wavelet transform), but as a superposition of the related *Schauder functions*  $e_0, e_{n,k}, k = 1, 2, \dots, 2^{n-1}, n = 1, 2, \dots$ , defined as follows

$$e_0(t) = t, \quad e_{n,k}(t) = \int_0^t \varphi_{n,k}(s) ds, \quad 0 \leq t \leq 1, \quad k = 1, 2, \dots, 2^{n-1}, n = 1, 2, \dots$$

For  $X \in C[0, 1]_0$ , we set

$$h_{n,k}(X) := X((k-1)/2^{n-1}) - \frac{1}{2}(X(k/2^{n-1}) + X((k-1)/2^{n-1}))$$

and

$$Y_0(X) = X(1), \quad Y_{n,k}(X) := 2^{(n+1)/2} \cdot h_{n,k}(X), \quad k = 1, 2, \dots, 2^{n-1}, n = 1, 2, \dots \quad (1.1)$$

$$\text{Let } X^N(t) = Y_0(X) \cdot t + \sum_{n=1}^N \sum_{k=1}^{2^{n-1}} Y_{n,k}(X) e_{n,k}(t).$$

We then have

**Lemma 1.1** (i)  $X^N$  is the linear interpolation of  $X$  on the  $N$ -th dyadic partition of  $[0, 1]$ . This implies

$$\lim_{N \rightarrow \infty} \sup_{t \in [0, 1]} |X^N(t) - X(t)| = 0.$$

(ii) Let  $\langle X \rangle_1^{2^N} := \sum_{k=1}^{2^N} (X(k/2^N) - X((k-1)/2^N))^2$  be the quadratic variation of  $X$  on the  $N$ -th dyadic partition. We then have

$$\langle X \rangle_1^{2^N} = \frac{1}{2^N} \left( Y_0^2 + \sum_{n=1}^N \sum_{k=1}^{2^{n-1}} (Y_{n,k})^2 \right).$$

(iii)  $X$  is absolutely continuous with  $X' \in L^2[0, 1]$ , i.e.  $X$  is in the Cameron–Martin space  $H$ , if and only if  $Y_0^2 + \sum_{n=1}^{\infty} \sum_{k=1}^{2^{n-1}} (Y_{n,k})^2 < \infty$ . In this case, we have

$$\|X\|_H^2 = \|X'\|_{L^2[0, 1]}^2 = Y_0^2 + \sum_{n=1}^{\infty} \sum_{k=1}^{2^{n-1}} (Y_{n,k})^2.$$

*Sketch of a proof.* (i) follows by induction on  $N$ ;

(ii) follows by induction on  $N$ .

(iii) follows by the representation of  $X'$  as a Fourier series with respect to the complete orthonormal system of Haar functions.  $\square$

Probability did not enter until now. Let now  $P$  be Wiener measure on  $C[0, 1]_0$ .

**Theorem 1.2**  $Y_0, Y_{n,k}, k = 1, 2, \dots, 2^{n-1}, n = 1, 2, \dots$  are iid random variables with distribution  $N(0, 1)$  under  $P$ .

*Sketch of a proof.* It is enough to show that  $Y_0, Y_{n,k}, k = 1, 2, \dots, 2^{n-1}, n = 1, 2, \dots$  are pairwise uncorrelated with distribution  $N(0, 1)$  under  $P$ .  $\square$

**Remark 1.3** Theorem 1.2 leads to an algorithm for the simulation of the Brownian path. The law of large numbers and Lemma 1.1(ii) imply the following:

**Corollary 1.4**  $\langle X \rangle_1^{2^N}$  converges  $P$ -a.s. and in  $L^2(P)$  to 1 if  $N \rightarrow \infty$ .

*Remark 1.5* This construction of Brownian motion goes back to Lévy and Ciesielski (see [13] for references), it can be carried out for every complete orthonormal system in  $L^2[0, 1]$ . Let  $\{\varphi_k | k \geq 1\}$  be a complete orthonormal system in  $L^2[0, 1]$ .

$$e_k(t) := \int_0^t \varphi_k(s) ds, \quad 0 \leq t \leq 1, \quad k = 1, 2, \dots$$

Let  $Y_k, k = 1, 2, \dots$  be iid with distribution  $N(0, 1)$  and set

$$X_t := \sum_{k=1}^{\infty} Y_k e_k(t) \quad 0 \leq t \leq 1.$$

Then  $(X_t)_{0 \leq t \leq 1}$  is a Brownian motion (see Itô and Nisio [11] for a proof).

Let  $I_0 := \{0, (n, k), k = 1, 2, \dots, 2^{n-1}, n = 1, 2, \dots\}$ . We interpret  $I_0$  as a binary tree:

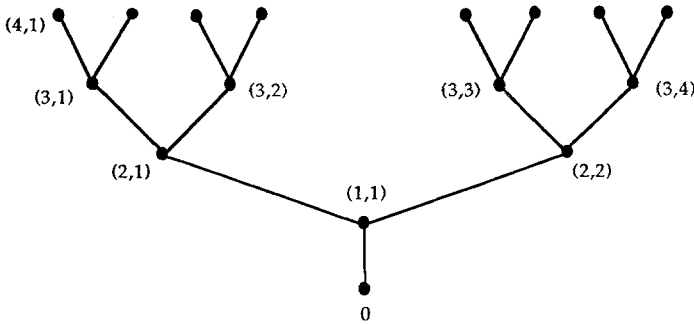


Fig. 1

To each function in  $C[0, 1]_0$  corresponds a set of coefficients  $Y_0, Y_{n,k}, k = 1, 2, \dots, 2^{n-1}, n = 1, 2, \dots$  according to the mapping from  $C[0, 1]_0$  to  $\mathbb{R}^{I_0}$  defined in (1.1). The converse, however, is not true: not each element of  $\mathbb{R}^{I_0}$  is a function in  $C[0, 1]_0$ , hence not each probability distribution on  $\mathbb{R}^{I_0}$  is a probability distribution on  $C[0, 1]_0$ .

**Lemma 1.6** Let  $Q$  be a probability distribution on  $\mathbb{R}^{I_0}$ ,  $M_n := \max_{k=1,2,\dots,2^{n-1}} |Y_{n,k}|$  and  $X^N$  as in Lemma 1.1. If

$$\sum_{n=1}^{\infty} Q[M_n > 2^{\alpha n}] < \infty \text{ for an } \alpha < 1/2 \tag{1.2}$$

then  $Q[X^m \text{ converges uniformly}] = 1$ , i.e.  $Q$  is a probability distribution on  $C[0, 1]_0$ .

*Proof.* We equip  $C[0, 1]_0$  with the supremum norm  $\|X\| = \sup_{t \in [0, 1]} |X(t)|$ . It is enough to show that

$$\sum_{n=1}^{\infty} Q[\|X^n - X^{n-1}\| > a_n] < \infty \quad (1.3)$$

for a sequence  $(a_n)_{n=1, 2, \dots} \in \mathbb{R}$  with  $\sum_{n=1}^{\infty} a_n < \infty$ .

The Borel–Cantelli lemma then implies that  $(X^n)_{n=1, 2, \dots}$  forms  $Q$ -a.s. a Cauchy sequence in  $C[0, 1]_0$ . To show (1.3), we note that

$$\|X^n - X^{n-1}\| \leq M_n \cdot 2^{-(n+1)/2}, \quad (1.4)$$

hence

$$Q[\|X^n - X^{n-1}\| > a_n] \leq Q[M_n > 2^{(n+1)/2} a_n].$$

If we set

$$a_n = 2^{-\beta n - (1/2)} \text{ with } \beta = (1/2) - \alpha > 0$$

and apply (1.2), the claim follows.  $\square$

## 2 Construction of a random sequence of probability distributions which converges to the distribution of the Brownian bridge

Let  $P$  be the distribution of the Brownian bridge. Then  $Y_0 = 0$   $P$ -a.s. and  $Y_{n,k}$ ,  $k = 1, 2, \dots, 2^{n-1}$ ,  $n = 1, 2, \dots$  are iid random variables under  $P$  with distribution  $N(0, 1)$ . We set

$$I := \{(n, k) | k = 1, 2, \dots, 2^{n-1}, n = 1, 2, \dots\}$$

and

$$\Omega = C[0, 1]_{0,0} := \{X \in C[0, 1] | X(1) = X(0) = 0\}, \quad X_t(\omega) := \omega(t).$$

We then have  $\Omega = C[0, 1]_{0,0} \subseteq \mathbb{R}^I$ . We denote the set of all probability distributions on  $\Omega$  by  $\mathcal{M}_1(\Omega)$ .

We consider the mapping  $T_0: \Omega \rightarrow \Omega$ , defined as

$$(T_0 X)_t = \sqrt{2}(X_{t/2} - tX_{1/2}) \quad 0 \leq t \leq 1.$$

$T_0$  corresponds to a shift to the left of the tree, i.e.  $Y_{n,k}(T_0 \omega) = Y_{n+1,k}(\omega)$ . In the same way, we define  $T_1: \Omega \rightarrow \Omega$  as

$$(T_1 X)_t := \sqrt{2}(X_{(t+1)/2} - (1-t)X_{1/2}) \quad 0 \leq t \leq 1.$$

$T_1$  corresponds to a shift to the right of the tree, i.e.  $Y_{n,k}(T_1 \omega) = Y_{n+1, 2^{n-1}+k}(\omega)$ .  $P$  is invariant under  $T_0$  and under  $T_1$ .

We now consider the bigger space  $\bar{\Omega}$ , defined as

$$\bar{\Omega} = \Omega \times \{0, 1\}^{\mathbb{N}}$$

$$\bar{\omega} = (\omega, \theta) \quad \theta \in \{0, 1\}^{\mathbb{N}}$$

$$\bar{P} := P \times \lambda,$$

where  $\lambda$  denotes product measure on  $\{0, 1\}^{\mathbb{N}}$  with  $\lambda[\theta_i = 0] = \lambda[\theta_i = 1] = 1/2$ . Let the shift  $\bar{T}$  on  $\bar{\Omega}$  be defined as

$$\begin{aligned} \bar{T}: \bar{\Omega} &\rightarrow \bar{\Omega} \\ (\omega, (\theta_1, \theta_2, \dots)) &\rightarrow (T_{\theta_1} \omega, (\theta_2, \theta_3, \dots)). \end{aligned}$$

$\bar{P}$  is invariant under  $\bar{T}$ . We can even show:

**Theorem 2.1**  $\bar{P}$  is ergodic with respect to  $\bar{T}$ .

*Proof.* Let  $\bar{\mathcal{F}}^* := \bigcap_n \sigma(\{\bar{T}^m, m > n\})$  be the tail-field on  $\bar{\Omega}$ .  $\bar{P}$  is a product measure on  $\bar{\Omega}$ , hence Kolmogorov's 0-1-law is satisfied:

$$\bar{P}[\bar{A}] = 0 \text{ or } \bar{P}[\bar{A}] = 1 \quad \text{if } \bar{A} \in \bar{\mathcal{F}}^*.$$

The  $\sigma$ -field  $\bar{S} := \{\bar{A} | \bar{T}^{-1} \bar{A} = \bar{A}\}$ , generated by the shift-invariant sets, is contained in  $\bar{\mathcal{F}}^*$ , hence we have  $\bar{P}[\bar{A}] = 0$  or  $\bar{P}[\bar{A}] = 1$ , if  $\bar{A} \in \bar{S}$ , i.e.  $\bar{P}$  is ergodic under  $\bar{T}$ .  $\square$

Let  $\delta_{T^k \bar{\omega}}$  denote Dirac measure on  $\bar{T}^k \bar{\omega}$  and define the probability distribution  $\bar{R}_n(\bar{\omega})$  by  $\bar{R}_n(\bar{\omega}) := (1/n) \sum_{k=1}^n \delta_{\bar{T}^k \bar{\omega}}$  (where  $\bar{T}^0$  denotes the identity). Let  $R_{n,\theta}(\omega)$  denote the marginal distribution of  $\bar{R}_n(\bar{\omega})$  on  $\Omega$  for fixed  $\theta$ :

$$R_{n,\theta}(\omega) = \frac{1}{n} \sum_{k=1}^n \delta_{(T_\theta)^{k-1} \omega}$$

where

$$\begin{aligned} (T_\theta)^k \omega &= T_{\theta_k} \circ \dots \circ T_{\theta_1} \omega, \quad k \geq 1, \\ (T_\theta)^0 \omega &= \omega. \end{aligned}$$

For each  $\theta \in \{0, 1\}^{\mathbb{N}}$ ,  $R_{n,\theta}$  is a random variable with values in  $\mathcal{M}_1(\Omega)$ .

**Theorem 2.2** For  $\lambda$ -a.e.  $\theta$  and  $P$ -a.e.  $\omega$ ,  $R_{n,\theta}(\omega)$  converges weakly to  $P$ .

*Proof.* Let  $f: \bar{\Omega} \rightarrow \mathbb{R}$  be measurable and bounded. With Birkhoff's ergodic theorem we get from Theorem 2.1:

$$\frac{1}{n} \sum_{k=0}^{n-1} f \circ \bar{T}^k \xrightarrow[n \rightarrow \infty]{} \int f d\bar{P} \quad \bar{P} - \text{a.s.},$$

i.e.

$$\int f dR_{n,\theta}(\omega) \rightarrow \int f d\bar{P} \quad \text{for } \lambda\text{-a.e. } \theta \text{ and } P\text{-a.e. } \omega.$$

Since the set of all bounded, continuous functions on  $\bar{\Omega}$  is countably generated, this implies

$$R_{n,\theta} \xrightarrow{w} P \quad \text{for } \lambda\text{-a.e. } \theta \text{ and } P\text{-a.e. } \omega. \quad \square$$

*Remark 2.3* In fact,  $(R_{n,\theta})$  converges weakly to  $P$  for each  $\theta \in \{0, 1\}^{\mathbb{N}}$  (this was shown by Ben Arous and Tamura in an unpublished paper).

*Remark 2.4* Theorem 2.2 says that we can reconstruct Wiener measure with an ‘‘infinitesimal piece’’ of a single ‘‘typical’’ path around a ‘‘typical’’ point of the unit

interval. This property characterizes fractals: the information about a fractal object is contained in an arbitrary small part of the object. Of course, we are dealing with random fractals: the invariance of  $\bar{P}$  under  $\bar{T}$  corresponds to a “self-similarity in distribution”:  $(\sqrt{2}(X_{t/2} - tX_{1/2}))_{0 \leq t \leq 1}$  and  $(\sqrt{2}(X_{(t+1)/2} - (1-t)X_{1/2}))_{0 \leq t \leq 1}$  have the same distribution under  $P$  as  $(X_t)_{0 \leq t \leq 1}$ . We get a deterministic fractal, if we set, for instance, all the coefficients  $Y_{n,k}$ ,  $k = 1, 2, \dots, 2^{n-1}$ ,  $n = 1, 2, \dots$  to the value 1, see Sect. 5, Example 5.3.

*Remark 2.5* Let  $Q \in \mathcal{M}_1(\Omega)$  and let  $\bar{Q} := Q \times \lambda$ . Then  $Q \ll P$  implies  $\bar{Q} \ll \bar{P}$ , hence  $R_{n,\theta}(\omega) \rightarrow P$   $\bar{Q}$ -a.s., i.e. for  $\lambda$ -a.e.  $\theta$ ,  $Q$ -a.e.  $\omega$ . This means  $Q$  cannot be reconstructed in this way, but we get  $P$  from a path which is “typical” for  $Q$ . The intuition behind this is the following: the drift of  $Q$  with respect to  $P$  is lost because of the iterated rescaling, see also Lemma 1.1(iii).

### 3 Generalization to probability distributions on $S^I$

Let  $S$  be a Polish space. We may now replace  $\mathbb{R}$  with  $S$  and consider  $\Omega = S^I$ . Let  $\mu$  be a probability distribution on  $S$  and  $P = \prod_{i \in I} \mu$  be the corresponding product measure on  $S^I$ .

We define  $T_0, T_1, \bar{Q}, \bar{P}, \bar{T}, \bar{R}_n$  and  $R_{n,\theta}$  as in Sect. 2. We equip  $S^I$  with the product topology and  $\mathcal{M}_1(S^I)$ , the set of probability distributions on  $S^I$ , with the topology of weak convergence. We then get, in the same way as Theorem 2.2:

**Theorem 3.1** For  $\lambda$ -a.e.  $\theta$  and  $P$ -a.e.  $\omega$ ,  $R_{n,\theta}(\omega)$  converges to  $P$  (in  $\mathcal{M}_1(S^I)$ ).

**Definition 3.2** We call  $Q$  stationary or self-similar, if  $Q$  is invariant under  $T_0$  and under  $T_1$ . Let  $\mathcal{M}_1^s$  denote the set of all stationary probability distributions on  $\Omega$ . We call  $Q \in \mathcal{M}_1^s$  ergodic, if  $\bar{Q} := Q \times \lambda$  is ergodic on  $\bar{\Omega}$  with respect to  $\bar{T}$ .

*Remark 3.3* If  $Q$  is ergodic,  $R_{n,\theta}(\omega)$  converges to  $Q$  for  $Q$ -a.e.  $\omega$  and  $\lambda$ -a.e.  $\theta$ .

### 4 Large deviations

In the following we investigate large deviations of the convergence of  $R_{n,\theta}$  to  $P$  in Theorem 3.1. Note that  $R_{n,\theta}$  would not correspond, on the lattice  $\mathbb{Z}^d$ , to the usual empirical field, but rather to the empirical distribution of a sequence of configurations we get shifting along the path of a (transient) random walk on the lattice (see also Remark 4.8). Ben Arous and Tamura got, for each fixed  $\theta$ , a large deviation principle for the distributions of  $R_{n,\theta}$ , where the rate function depends heavily on  $\theta$  (see [3]). We are interested in estimates which do not depend on  $\theta$ . In particular, our rate function will be finite only on probability distributions which are stationary, i.e. invariant under all shifts  $T_\theta$ . There is, however, no large deviation principle holding “uniformly” in  $\theta$ : we can get the bounds in Theorem 4.1 only for  $\lambda$ -a.e.  $\theta$ . To understand why, look at the following example:

Let  $I = \mathbb{R}$  and  $A := \{Q \mid Q[Y_{2,2} > Y_{1,1}] = 1\}$ . If we shift only to the left, i.e. take  $\theta = (0, 0, 0, \dots)$ , we have  $P[R_{n,\theta} \in A] = P[Y_{2,2} > Y_{1,1}]^n$ , hence  $(1/n) \log P[R_{n,\theta} \in A] = \log P[Y_{2,2} > Y_{1,1}]$ . If we shift only to the right, i.e. take  $\theta = (1, 1, 1, \dots)$ , we get  $P[R_{n,\theta} \in A] = P[Y_{1,1} < Y_{2,2} < Y_{3,4} < \dots < Y_{n, 2^{n-1}}]$ , hence  $(1/n) \log P[R_{n,\theta} \in A] \rightarrow -\infty$ .

Also, there is no “global” rate function concentrated on  $\mathcal{M}_1^s(\Omega)$ . This will follow from Theorem 4.1. We give an illustrative example:

Let  $I = \mathbb{R}$  and  $P = \prod_{i \in I} N(0, 1)$  as in Sect. 2. Consider  $A_m : \{Q | E_Q[Y_{m,1}] = E_Q[Y_{m,2}] = \dots = E_Q[Y_{m,2^{m-1}}] \geq b\}$  ( $m = 1, 2, \dots$ ) where  $b > 0$ . Since  $P$  is a product measure, we have, for all  $\theta$ ,

$$P[R_{n,\theta} \in A_m] = P[\bar{Y}_n \geq b]^{2^{m-1}},$$

where  $\bar{Y}_n$  is the arithmetic mean of  $Y_{1,1}, Y_{2,1}, Y_{3,1}, \dots, Y_{n,1}$ . So we get

$$\lim_n \frac{1}{n} \log P[R_{n,\theta} \in A_m] = -2^{m-1} \lambda(b)$$

for all  $\theta$ , where  $\lambda(x) = x^2/2$  is the rate function for the large deviations of the arithmetic mean of iid random variables with distribution  $N(0, 1)$ . On the other hand,  $A_m \cap \mathcal{M}_1^s(\Omega) = A_1 \cap \mathcal{M}_1^s(\Omega)$  for each  $m$ .

We establish a principle of large deviations for the finite-dimensional marginal distributions of  $R_{n,\theta}$  on  $S^J$  ( $J \subseteq I, J$  finite) under  $P$ . Recall  $P$  is, in this section, a product measure on  $S^I$ . Let us begin with some notation: we denote the mapping of the index set  $I$  on itself, which corresponds to  $T_0$ , again by  $T_0$ , i.e.  $T_0(n, k) = (n + 1, k)$  and, in the same way,  $T_1(n, k) = (n + 1, 2^{n-1} + k)$  ( $k = 1, 2, \dots, 2^{n-1}, n = 1, 2, \dots$ ). Let  $J \subseteq I$  and let  $T_{\theta_0}$  denote the identity. Let  $F_k(\theta, J) := \bigcup_{\ell=0}^{k-1} T_{\theta_\ell} \circ \dots \circ T_{\theta_0} J$  be the set of coordinates, generated by  $J$  after  $k - 1$  shifts according to  $\theta$ , ( $k \geq 1$ ),  $F_0(J) := J$ . Let  $\mathcal{F}_k(\theta, J) := \sigma(\{Y_i | i \in F_k(\theta, J)\})$  be the corresponding  $\sigma$ -field. We write  $(T_\theta)^\ell$  instead of  $T_{\theta_\ell} \circ \dots \circ T_{\theta_0}$ , hence  $F_k(\theta, J) = \bigcup_{\ell=0}^{k-1} (T_\theta)^\ell J$ , ( $k \geq 1$ ). Let  $\mathcal{M}_1(S^J)$  denote the set of all probability distributions on  $S^J$  as before. We now consider subsets  $A_J$  of  $\mathcal{M}_1(S^J)$ , which are characterized in the following way: let  $J \subseteq I, J$  finite,  $B_J \subseteq \mathcal{M}_1(S^J)$  measurable and

$$A_J := \{Q | Q|_{\mathcal{F}_0(J)} \in B_J\} \quad (4.1)$$

i.e. if  $Q$  is in  $A_J$  or not depends only on the finite-dimensional marginal of  $Q$  on  $\mathcal{F}_0(J)$ . We then have:  $A_J$  is open in  $\mathcal{M}_1(\Omega) \Leftrightarrow B_J$  is open in  $\mathcal{M}_1(S^J)$ .

Let the *relative entropy*  $H(Q|P)|_{\mathcal{F}}$  of  $Q$  with respect to  $P$  on the  $\sigma$ -field  $\mathcal{F}$  be defined as  $E_Q \left[ \log \frac{dQ}{dP} \Big|_{\mathcal{F}} \right]$ , if  $Q \ll P$  on  $\mathcal{F}$ , and  $= +\infty$  else. Now we can state the following large deviation principle:

**Theorem 4.1** *Let  $J \subseteq I, J$  finite. Then there is a function  $I_J : \mathcal{M}_1(\Omega) \rightarrow [0, \infty]$ , such that the following holds for all  $A_J$  of the form in (4.1):*

$$\begin{aligned} A_J \text{ open} &\Rightarrow \overline{\lim}_n \frac{1}{n} \log P[R_{n,\theta} \in A_J] \geq - \inf_{Q \in A_J} I_J(Q) \\ A_J \text{ closed} &\Rightarrow \overline{\lim}_n \frac{1}{n} \log P[R_{n,\theta} \in A_J] \leq - \inf_{Q \in A_J} I_J(Q) \end{aligned}$$

for  $\lambda$ -a.e.  $\theta$ . Further,  $I_J : \mathcal{M}_1(\Omega) \rightarrow [0, \infty]$  is lower semicontinuous,  $I_J(Q) = +\infty$  if  $Q \notin \mathcal{M}_1^s$  and we have for  $Q \in \mathcal{M}_1^s$ :

$$I_J(Q) = \lim_n \frac{1}{n} \int H(Q|P)|_{\mathcal{F}_n(\theta, J)} \lambda(d\theta).$$



*Proof.* We fix  $J$  and write  $\mathcal{F}_0, \mathcal{F}_n(\theta), A$  instead of  $\mathcal{F}_0(\theta, J), \mathcal{F}_n(\theta, J), A_J$ . The basic idea of the proof is to represent  $R_{n,\theta}|_{\mathcal{F}_0}$  as the empirical distribution of a Markov chain of order  $m$ . We will give an explicit proof of the lower bound and refer to a general theorem in the proof of the upper bound.

First we investigate the behaviour of the Radon–Nikodym derivatives of  $Q$  with respect to  $P$  on the  $\sigma$ -fields  $\mathcal{F}_n(\theta)$ : in fact, we need to know only the growths of the relative entropies  $H(Q|P)|_{\mathcal{F}_n(\theta)}$  of  $Q$  with respect to  $P$  on the  $\sigma$ -fields  $\mathcal{F}_n(\theta)$ .

**Theorem 4.2** *For each  $Q \in \mathcal{M}_1^s$ ,  $(1/n)H(Q|P)|_{\mathcal{F}_n(\theta)}$  converges to  $I_J(Q)$  for  $\lambda$ -a.e.  $\theta$ , where*

$$I_J(Q) = \lim_n \frac{1}{n} \int H(Q|P)|_{\mathcal{F}_n(\theta)} \lambda(d\theta) \in [0, \infty]$$

*Further,  $I_J(\cdot)$  is lower semicontinuous and affine.*

*Remark 4.3* If  $Q$  is ergodic, we have for  $\lambda$ -a.e.  $\theta$ :  $(1/n) \log \frac{dQ}{dP}|_{\mathcal{F}_n(\theta)}$  converges  $Q$ -a.s. to  $I_J(Q)$ .

*Proof of Theorem 4.2* Let  $\bar{Q} = Q \times \lambda, \bar{P} = P \times \lambda$ . We make use of a theorem of A. Barron .

**Theorem 4.4** *Let  $(X_n)_{n=1,2,\dots}$  be a stationary process with values in a Standard Borel space  $E$ . Let  $\bar{Q}$  be the distribution of  $(X_n)$ , and  $P \in \mathcal{M}_1(E^{\mathbb{N}})$  be a “reference measure”:  $P$  is stationary and Markov of order  $m$  (i.e.  $P[A|X_{n-1}, \dots, X_1] = P[A|X_{n-1}, \dots, X_{n-m}]$  for all  $\sigma(X_n, X_{n+1}, \dots)$  measurable sets  $A$  and  $n > m$ ). Define the  $\sigma$ -fields  $A_n := \sigma(X_1, X_2, \dots, X_n)$ . Then the specific relative entropy  $h(Q|P)$  of  $Q$  with respect to  $P$  exists:*

$$h(Q|P) = \lim_n \frac{1}{n} H(Q|P)|_{\mathcal{A}_n} \in [0, \infty] .$$

See Barron [1, Theorem 1] for the proof.

Note that  $H(Q|P)|_{\mathcal{A}_n}$  is increasing, so  $h(Q|P) = \infty$  if there is an  $n$  such that  $H(Q|P)|_{\mathcal{A}_n} = \infty$ .

We apply Theorem 4.4 with  $\bar{Q}, \bar{P}, E := \{(Y_1, \dots, Y_\ell) | Y_i \in S\} \times \{0, 1\}$ , where  $(1, 2, \dots, \ell)$  is an enumeration of  $J$ .

$\bar{P}|_{\sigma(J, T_\theta J, T_\theta^2 J, \dots) \times \mathcal{B}}$  can be identified with a stationary measure, Markov of order  $m$  on  $E^{\mathbb{N}}$ , where  $X_n$  consists of the  $\ell$ -tuple of random variables  $Y_i, i \in T_\theta^{n-1} J$  and  $\theta_n$ .

$$m := \max \{n | \exists k \text{ with } (n, k) \in J\} - 1. \quad (4.2)$$

(Since the sets  $T_\theta^n J, n = 1, 2, \dots$  are not disjoint in general,  $\bar{P}|_{\sigma(J, T_\theta J, T_\theta^2 J, \dots) \times \mathcal{B}}$  is in general not a product measure on  $E^{\mathbb{N}}$ ). In the same way, we can identify a stationary and ergodic  $\bar{Q}$  with a stationary and ergodic probability distribution on  $E^{\mathbb{N}}$ , respectively. Hence we get from Theorem 4.4

$$\exists \bar{I}_J(\bar{Q}) = \lim_n \frac{1}{n} H(\bar{Q}|\bar{P})|_{\sigma(X_1, \dots, X_n)} .$$

If we set  $I_J(Q) := \overline{I_J(Q)}$ , we get

$$I_J(Q) = \lim_n \frac{1}{n} \int H(Q|P)|_{\mathcal{F}_n(\theta)} \lambda(d\theta). \quad (4.3)$$

It remains to show that  $I_J$  is lower semicontinuous and affine; for this, we refer to [9].  $\square$

*Remark 4.5*  $Q \in \mathcal{M}_1^s$  and  $I_J(Q) = 0$  imply  $Q = P$ .

The next step in the proof of the lower bound is to show that the set of ergodic probability distributions is dense in  $\mathcal{M}_1^s$ .

**Lemma 4.6** *Let  $Q \in \mathcal{M}_1^s$ . Then there is a sequence of ergodic probability distributions  $(Q_n)_{n=1, 2, \dots}$  such that  $Q_n \xrightarrow{w} Q$ , and  $I_J(Q_n) \rightarrow I_J(Q)$  for  $n \rightarrow \infty$ .*

The proof is similar to the proof of Lemma 4.8 in Föllmer [7].

Note that Theorem 4.2, Lemma 4.6 and Remark 3.3 can be used to prove the lower bound with a standard argument (see [5, p. 76]).

For each  $\theta$ ,  $R_{n,\theta}|_{\mathcal{F}_0}$  is the  $n$ -th empirical distribution of a Markov chain of order  $m$ . This chain is, in the terminology of Deuschel et al. [6]  $R$ -mixing with  $M = 1$  if  $R \geq m$  (see (4.2)), since  $\sigma(\{T_\theta^{k\omega}, 0 \leq k \leq r\})$  and  $\sigma(\{T_\theta^{k\omega}, k \geq r+m\})$  are independent. A general theorem about uniform large deviations (see [6, p. 91]) implies, together with the contraction principle, that for each  $\theta$ , the distributions of  $R_{n,\theta}|_{\mathcal{F}_0}$ ,  $n \geq 1$ , satisfy a large deviation principle with rate function  $h_{\theta,J}: \mathcal{M}_1(\mathcal{F}_0) \rightarrow [0, \infty]$  and  $h_{\theta,J}(v) = \inf \{I_{\theta,J}(Q) | Q \in \mathcal{M}_1^s(\theta), Q|_{\mathcal{F}_0} = v\}$  with  $I_{\theta,J}(Q) = \lim_n (1/n) H(Q|P)|_{\mathcal{F}_n(\theta)}$ . (Here,  $\mathcal{M}_1^s(\theta)$  denotes the set of probability distributions which are invariant under  $T_\theta$ ). It remains to identify this rate function with the rate function in Theorem 4.1. Theorem 4.2 implies that for  $\lambda$ -a.e.  $\theta$ ,  $\inf \{I_{\theta,J}(Q) | Q \in \mathcal{M}_1^s(\theta), Q|_{\mathcal{F}_0} = v\} = \inf \{I_{\theta,J}(Q) | Q \in \mathcal{M}_1^s, Q|_{\mathcal{F}_0} = v\}$  and  $\lim_n (1/n) H(Q|P)|_{\mathcal{F}_n(\theta)} = I_J(Q)$ , so  $I_{\theta,J}(Q) = I_J(Q)$  for  $\lambda$ -a.e.  $\theta$ .  $\square$

Note that for each  $L \geq 0$ ,  $\{Q|_{\mathcal{F}_0} | I_J(Q) \leq L\}$  is a compact subset of  $\mathcal{M}_1(S^J, \mathcal{F}_0)$ , but  $\{Q | I_J(Q) \leq L\}$  is in general not compact in  $\mathcal{M}_1(\Omega)$ .

Of course, the arguments in [6] are much more general than our situation requires. For a direct proof of Theorem 4.1, we refer to [9].

Theorem 4.1 says that we have to minimize the rate functional  $I_J$  over the set of probability distributions  $\mathcal{M}_1^s(S^I)$ . It is therefore natural to ask about the properties of probability distributions in  $\mathcal{M}_1^s(S^I)$ . We omitted  $Y_0$ , but we can extend any stationary measure on  $S^I$  to a stationary measure on  $S^{I_0}$ . Let  $P = \prod_{i \in I_0} \mu \in \mathcal{M}_1^s(S^{I_0})$  denote a product measure and consider  $\mathcal{M}_1^s = \mathcal{M}_1^s(S^{I_0})$ . Probability distributions in  $\mathcal{M}_1^s$  are typically singular with respect to  $P$ . More precisely,  $Q \in \mathcal{M}_1^s$ ,  $Q \ll P \Rightarrow Q = P$ . In particular,  $Q \in \mathcal{M}_1^s$  has infinite relative entropy with respect to  $P$ , if  $Q \neq P$ . We can show, though, that a specific relative entropy with respect to  $P$  exists for each  $Q \in \mathcal{M}_1^s$ . Consider the  $\sigma$ -fields  $\mathcal{F}_{2^n} = \sigma(\{Y_0, Y_{m,k} | m \leq n\})$ .

**Lemma 4.7** Every  $Q \in \mathcal{M}_1^s$  has a specific relative entropy  $h(Q|P)$  with respect to  $P$ :

$$\begin{aligned} h(Q|P) &= \lim_n \frac{1}{2^n} H(Q|P)|_{\mathcal{F}_{2^n}} \\ &= \sup_n \frac{1}{2^n} H(Q|P)|_{\mathcal{F}_{2^n}} \in [0, \infty] . \end{aligned}$$

In particular:  $Q \in \mathcal{M}_1^s$ ,  $h(Q|P) = 0 \Rightarrow Q = P$ .

Further,  $h(\cdot|P)$  is affine on  $\mathcal{M}_1^s(\Omega)$ .

*Sketch of a proof.* Let  $\mathcal{V}$  be the set containing all subsets of  $I_0$ . Then the function  $f: \mathcal{V} \rightarrow \mathbb{R}$ ,  $f(V) = H(Q|P)|_{\sigma(\{Y_i | i \in V\})}$  is a *superadditive set function*, i.e.  $f(V \cup W) \geq f(V) + f(W)$  for disjoint sets  $V, W \in \mathcal{V}$ . Here we made use of the product structure of  $P$ . The rest of the proof is left to the reader (see also Georgii, [10, Chap. 15, Sect. 2], for a general argument).  $\square$

**Remark 4.8** Look at

$$R_n(\omega) := \frac{1}{2^n} \sum_{\theta \in \{0, 1\}^n} \sum_{k=1}^n \delta_{(T_\theta)^k \omega} .$$

$R_n(\omega)$  is the analogon to the usual empirical field on the tree. Then the same arguments as in the proof of Theorem 2.2 show that  $R_n(\omega) \rightarrow P$  for  $P$ -a.e.  $\omega$ . The distributions of  $R_n$ ,  $n \geq 1$ , satisfy a large deviation principle with good rate function  $I$  where  $I(Q) = h(Q|P)$  for  $Q \in \mathcal{M}_1^s$ ,  $I(Q) = +\infty$  else. We don't give a proof here, since it consists merely in carrying over arguments in [6] or [8] from the lattice to the tree structure. (The arguments in [6] or [8] can here be simplified of course, since we treat the particularly nice case of a product measure  $P$ .)

## 5 Examples

### 5.1 Stationary probability distributions on $\mathbb{R}^{I_0}$

Let  $P = \prod_{i \in I_0} N(0, 1) \in \mathcal{M}_1^s(\mathbb{R}^{I_0})$  denote Wiener measure. Can we identify a self-similar probability distribution on  $\mathbb{R}^{I_0}$  with a probability distribution on  $C[0, 1]_0$ ? Notice this is not clear *a priori*. In Lemma 5.1 below we give a condition on  $Q$  which guarantees that the support of  $Q$  is contained in  $C[0, 1]_0$ . Our conjecture is, however, that this holds true for every  $Q \in \mathcal{M}_1^s(\mathbb{R}^{I_0})$ . Because we deal with stationary probability distributions, condition (5.1) below involves only the one-dimensional marginal distribution of  $Q$ .

**Lemma 5.1** Let  $Q \in \mathcal{M}_1^s = \mathcal{M}_1^s(\mathbb{R}^{I_0})$  and

$$\sum_{n=1}^{\infty} 2^{n-1} Q[|Y_0| \geq 2^{2^n}] < \infty \quad \text{for an } \alpha < 1/2 . \quad (5.1)$$

Then  $Q$  is a probability distribution on  $C[0, 1]_0$ .

*Proof.* We can replace (1.2) in Lemma 1.6 with (5.1):

$$\begin{aligned} Q[M_n \geq 2^{2^n}] &\leq Q\left[\max_{k=1, 2, \dots, 2^{n-2}} |Y_{n,k}| \geq 2^{2^n}\right] + Q\left[\max_{k=2^{n-2}+1, \dots, 2^{n-1}} |Y_{n,k}| \geq 2^{2^n}\right] \\ &= 2Q[M_{n-1} \geq 2^{2^n}] , \end{aligned}$$

since  $Q$  is stationary. By iteration, we conclude  $Q[M_n \geq 2^{2^n}] \leq 2^{n-1}Q[|Y_0| \geq 2^{2^n}]$ .  $\square$

Let  $\nu$  be the distribution of  $Y_0$  under  $Q$ . Sufficient for (5.1) to hold is, for instance,  $\int |x|^{2+\varepsilon} d\nu < \infty$  for an  $\varepsilon > 0$  or  $\sup_{t \in \mathbb{R}} \left| \frac{d\nu}{d\mu}(t) \right| \leq C$ , where  $\mu = N(0, 1)$ .

If the support of  $Q$  is contained in  $C[0, 1]_0$ , we get the coordinate process  $(X_t)_{0 \leq t \leq 1}$  via

$$X_t(\omega) = Y_0(\omega) \cdot t + \sum_{n=1}^{\infty} \sum_{k=1}^{2^{n-1}} Y_{n,k}(\omega) e_{n,k}(t) \quad (0 \leq t \leq 1) \tag{5.2}$$

as in Sect. 1. In this case, we can write the  $\sigma$ -field  $\mathcal{F}_{2^n}$  in Lemma 4.7 as  $\mathcal{F}_{2^n} = \sigma(\{X_{k \cdot 2^{-n}} | k = 0, 1, \dots, 2^n\})$ , and we have an interpretation of the specific relative entropy  $h(Q|P)$  as a limit of entropies on the dyadic partitions of the unit interval.

The simplest case of a stationary probability distribution is, of course, a product measure  $Q = \prod_{i \in I_0} \nu$ . What can we then say about  $(X_t)_{0 \leq t \leq 1}$ ?

**Lemma 5.2** *Let  $Q = \prod_{i \in I_0} \nu$  and assume  $Q$  satisfies (5.1).*

(i) *Let  $\int x^2 d\nu < \infty$ . Then  $E_Q[X_t X_s] = \int x^2 d\nu \cdot (t \wedge s)$ . Further,  $(X_t)_{0 \leq t \leq 1}$  has quadratic variation  $\langle X_t \rangle_{0 \leq t \leq 1}$  (along the sequence of dyadic partitions of  $[0, 1]$ ) and*

$$\langle X \rangle_t = (\int x^2 d\nu) \cdot t \quad (0 \leq t \leq 1), \text{ } Q\text{-a.s.}$$

(ii) *Let  $\int x^4 d\nu < \infty$ . Then  $(X_t)_{0 \leq t \leq 1}$  is locally Hölder-continuous with exponent  $\gamma$ , for each  $\gamma < 1/2$ .*

For the proof, we refer to [9].

In general,  $(X_t)_{0 \leq t \leq 1}$  is not a Markov process under  $Q$ . We can state the following “weakened Markov property”: the conditional distribution of  $\{X_t | t \in ](k-1)/2^n, k/2^n[ \}$ , given  $\{X_t | t \in [0, (k-1)/2^n] \cup [k/2^n, 1]\}$ , depends only on  $X_{(k-1)/2^n}$  and  $X_{k/2^n}$ ,  $k = 1, 2, \dots, 2^n$ ,  $n = 1, 2, \dots$

*Example 5.3* Let  $\nu = N(a, 1)$  and  $Q = \prod_{i \in I} \nu$ . Of course, condition (5.1) in Lemma 5.1 is satisfied.  $Q$  is the distribution of  $(B_t + a \cdot g(t))_{0 \leq t \leq 1}$ , where  $(B_t)_{0 \leq t \leq 1}$  is

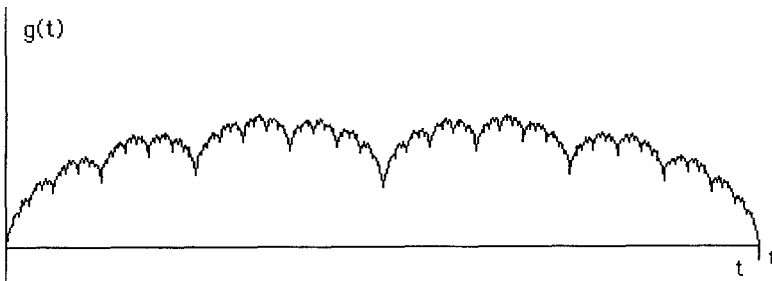


Fig. 2

a Brownian bridge and  $g$  the self-similar function  $g(t) = \sum_{n=1}^{\infty} \sum_{k=1}^{2^{n-1}} e_{n,k}(t)$  ( $0 \leq t \leq 1$ ).  $g$  is the profile of the fractal “mount Takagi” (see [14]).

### 5.2 Brownian sheet and other examples with $S = C[0, 1]$

Except Brownian motion (multiplied with constants), real-valued diffusions resp. their bridges are not self-similar in our sense, i.e. invariant under  $T_0$  and under  $T_1$ . But if we allow the coefficients  $Y_0, Y_{n,k}, k = 1, 2, \dots, 2^{n-1}, n = 1, 2, \dots$  to have values in a function space, we can describe well-known, “smooth” objects like the  $C[0, 1]$ -valued Ornstein–Uhlenbeck process. Take  $S = C[0, 1]$  and  $P^\infty = \prod_{t \in I_0} P \in \mathcal{M}_1^s(S^{I_0})$ , where  $P$  denotes Wiener measure on  $C[0, 1]$ . Set

$$X(t, \tau) = Y_0(\tau) \cdot t + \sum_{n=1}^{\infty} \sum_{k=1}^{2^{n-1}} Y_{n,k}(\tau) e_{n,k}(t) \quad (0 \leq t, \tau \leq 1). \quad (5.3)$$

$(X(t, \cdot))_{0 \leq t \leq 1}$  is then a  $C[0, 1]$ -valued Brownian motion under  $P^\infty$ : For  $t_1, t_2, t_3, t_4$  with  $0 \leq t_1 < t_2 \leq t_3 \leq t_4 \leq 1$  the increments  $X(t_2, \cdot) - X(t_1, \cdot)$ ,  $X(t_4, \cdot) - X(t_3, \cdot)$  are independent and  $1/(\sqrt{t_2 - t_1})(X(t_2, \cdot) - X(t_1, \cdot))$  has distribution  $P$ . In the same way,  $(X(\cdot, \tau))_{0 \leq \tau \leq 1}$  is a  $C[0, 1]$ -valued Brownian motion under  $P^\infty$ . We call  $P^\infty$  infinite-dimensional Wiener measure or the distribution of “Brownian sheet”. Let us replace  $P^\infty$  with another Gaussian product measure  $Q^\infty = \prod_{t \in I_0} Q \in \mathcal{M}_S(S^{I_0})$ .

**Lemma 5.4** *Let  $Q$  be a Gaussian probability distribution on  $C[0, 1]$  with  $E_Q[Y_0(\tau)] = 0$  ( $0 \leq \tau \leq 1$ ) and set*

$$X(t, \tau) = Y_0(\tau) \cdot t + \sum_{n=1}^{\infty} \sum_{k=1}^{2^{n-1}} Y_{n,k}(\tau) e_{n,k}(t) \quad (0 \leq t, \tau \leq 1)$$

where  $Y_0, Y_{n,k}, k = 1, 2, \dots, 2^{n-1}, n = 1, 2, \dots$  are independent with distribution  $Q$ . We then have

- (i) For each  $\tau \in [0, 1]$ ,  $(X(t, \tau))_{0 \leq t \leq 1}$  is a Brownian motion with variance  $E_Q[Y_0(\tau)^2]$ .
- (ii) The  $C[0, 1]$ -valued process  $(X(t, \cdot))_{0 \leq t \leq 1}$  has independent increments, and  $1/(\sqrt{t_2 - t_1})(X(t_2, \cdot) - X(t_1, \cdot))$  has distribution  $Q$  ( $0 \leq t_1 < t_2 \leq 1$ ).

The proof is easy: we refer to [9].

**Example 5.5** Let  $Q \in \mathcal{M}_1(C[0, 1])$  be the distribution of an Ornstein–Uhlenbeck process starting in 0, i.e. the distribution of  $(Z_t)_{0 \leq t \leq 1}$ , where  $(Z_t)$  solves the stochastic differential equation

$$dZ_t = dW_t - Z_t dt$$

and  $(W_u)_{0 \leq u \leq 1}$  is a Brownian motion under  $Q$ . Then the  $C[0, 1]$ -valued process  $(X(\cdot, \tau))_{0 \leq \tau \leq 1}$  is a  $C[0, 1]$ -valued Ornstein–Uhlenbeck process under  $Q^\infty$ , i.e.  $X_\tau := X(\cdot, \tau)$  solves the (“infinite-dimensional”) stochastic differential equation

$$dX_\tau = dW_\tau - X_\tau d\tau$$

where  $W_\tau := W(\cdot, \tau)$  is a  $C[0, 1]$ -valued Brownian motion under  $Q^\infty$ . We can describe  $Q^\infty$  with Theorem 4.1 as a “large deviation” of  $P^\infty$ , i.e. as the solution of a variational problem where we have to minimize the rate function in Theorem 4.1 over a certain subset of  $\mathcal{M}_1(C[0, 1]^{I_0})$ . More precisely, set

$$A := \{\bar{R} \in \mathcal{M}_1(C[0, 1]^{I_0}) \mid \bar{R}|_{\mathcal{F}_0} \in B\}$$

where

$$B := \{R \mid \int X_t^2 dR \leq 1 - e^{-t}, 0 \leq t \leq 1\},$$

$B \subseteq \mathcal{M}_1(C[0, 1])$ . Here  $A$  is of the form in (4.1), i.e.  $\bar{R} \in A$  iff the one-dimensional marginal distribution of  $\bar{R}$  is in  $B$ . The rate function  $I_J(\cdot)$  is here the specific relative entropy  $h(\cdot \mid P^\infty)$ . Since  $Q$  minimizes the relative entropy  $H(\cdot \mid P)$  over  $B$  (this is shown in [9]),  $Q^\infty$  minimizes  $h(\cdot \mid P^\infty)$  over  $A$ . In this way, we may see the  $C[0, 1]$ -valued Ornstein–Uhlenbeck process as a large deviation of Brownian sheet.

*Acknowledgement.* This work is part of a Ph.D. Thesis written at Bonn University. I would like to thank my adviser, Hans Föllmer, and also Michael Röckner for many helpful discussions.

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